PSY 503: Foundations of Psychological Methods

Lecture 19: Heterogeneity and Experimental Regression Adjustments

Robin Gomila

Princeton

November 11, 2020

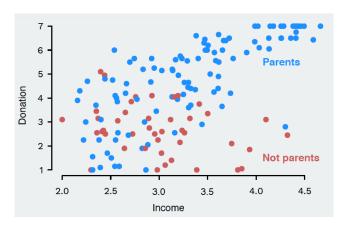
Heterogeneity

- ullet So far, we have considered cases in which eta_1 was constant across subgroups
- What happens if it's not?
- Let's open R Studio!

Why interaction terms?

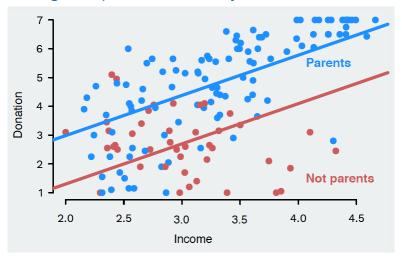
- Interaction terms will allow you to let the slope on one variable vary as a function of another variable
- Let's explore a different hypothetical dataset describing the relationship between income and donations conditional on having children

Let's see the data



• Does it look like there's heterogeneity? What happens if we simply control for "parents" additively?

Controlling for "parent" additively



• The regression is a poor fit for non-parents. Can we allow for different slopes for each group?

Interactions with a binary variable

- Let X_2 be binary
- In this case, $X_2 = 1$ for parents
- We can add another covariate to the baseline model that allows the effect of income to vary by parental status
- This covariate is called an interaction term and it is the product of the two marginal variables of interest: $income \times parent$
- Here is the model with the interaction term:

$$\widehat{Y} = \widehat{\beta}_0 + \widehat{\beta}_1 X_1 + \widehat{\beta}_2 X_2 + \widehat{\beta}_3 X_1 X_2$$

Two lines in one regression

$$\widehat{Y}_i = \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} + \widehat{\beta}_2 X_{2_i} + \widehat{\beta}_3 X_{1_i} X_{2_i}$$

- ullet How can we interpret this model? We can plug in the two possible values of X_{2i}
- When $X_{2_i} = 0$:

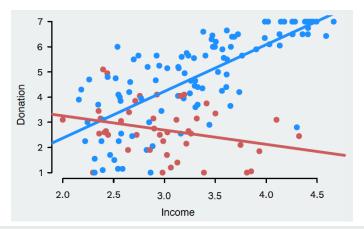
$$\begin{split} \widehat{Y}_i &= \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} + \widehat{\beta}_2 X_{2_i} + \widehat{\beta}_3 X_{1_i} X_{2_i} \\ &= \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} + \widehat{\beta}_2 \times 0 + \widehat{\beta}_3 X_{1_i} \times 0 \\ &= \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} \end{split}$$

• When $X_{2_i} = 1$:

$$\begin{split} \widehat{Y}_i &= \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} + \widehat{\beta}_2 X_{2_i} + \widehat{\beta}_3 X_{1_i} X_{2_i} \\ &= \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} + \widehat{\beta}_2 \times 1 + \widehat{\beta}_3 X_{1_i} \times 1 \\ &= (\widehat{\beta}_0 + \widehat{\beta}_2) + (\widehat{\beta}_1 + \widehat{\beta}_3) X_{1_i} \end{split}$$

Example interpretation of the coefficients

	Intercept for X_{2i}	Slope for X_{2i}
$X_{2i} = 0$	\widehat{eta}_{0}	\widehat{eta}_1
$X_{2i} = 1$	$\widehat{eta}_0 \ \widehat{eta}_0 + \widehat{eta}_2$	$\widehat{eta}_1 + \widehat{eta}_3$

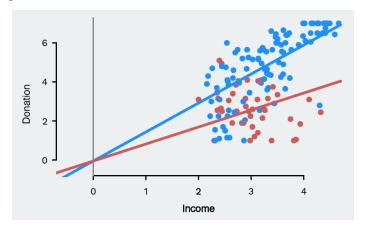


General interpretation of the coefficients

- ullet \widehat{eta}_0 : average value of Y_i when both X_{1_i} and X_{2_i} are equal to 0
- $\widehat{\beta}_1$: a one-unit change in X_{1_i} is associated with a $\widehat{\beta}_1$ -unit change in Y_i when $X_{2_i}=0$
- $\widehat{\beta}_2$: average difference in Y_i between $X_{2_i}=1$ group and $X_{2_i}=0$ group when $X_{1_i}=0$
- $\widehat{\beta}_3$: change in the effect of X_{1_i} on Y_i between $X_{2_i}=1$ group and $X_{2_i}=0$ group

Lower order terms

- Principle of Marginality: Always include the lower order terms!
- Imagine that we ommitted the lower order term:



Omitting lower order terms

$$\widehat{Y}_i = \widehat{\beta}_0 + \widehat{\beta}_1 X_{1_i} + 0 \times X_{2_i} + \widehat{\beta}_3 X_{1_i} X_{2_i}$$

- This model does not allow for a difference between parents and non-parents when income is 0
 - This distorts slope estimates
 - Very rarely justified
 - Yet, for some reasons, people do it...

A note on interactions with non-binary discrete and continuous variables

- Same principle!
- ullet Plug in values in the equation to get the marginal effect of X_{1_i} conditional on specific values of X_{i_2}

Summary for interactions

- Do not omit lower order terms because it usually imposes unrealistic restrictions
- Do not interpret the coefficients on the lower order terms as marginal effects
 - They give the marginal effect only for the case where the other variable is equal to zero
- The p-value on the interaction term can be used as a test against the null of no interaction, but significant tests for the lower order term rarely make sense

Polynomials

Polynomial terms

- Polynomial terms are a special case of the continuous variable interactions
 - where $X_1 = X_2$

$$\begin{split} \widehat{Y}_{i} &= \widehat{\beta}_{0} + \widehat{\beta}_{1}X_{1_{i}} + \widehat{\beta}_{2}X_{1_{i}} + \widehat{\beta}_{3}X_{1_{i}}X_{1_{i}} \\ &= \widehat{\beta}_{0} + (\widehat{\beta}_{1} + \widehat{\beta}_{2})X_{1_{i}} + \widehat{\beta}_{3}X_{1_{i}}X_{1_{i}} \\ &= \widehat{\beta}_{0} + (\widehat{\beta}_{1} + \widehat{\beta}_{2})X_{1_{i}} + \widehat{\beta}_{3}X_{1_{i}}^{2} \end{split}$$

• This is called a **second order polynomial** in X_1 , which we generally write:

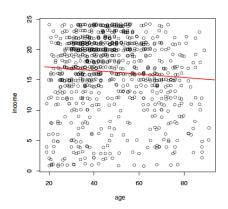
$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + u_i$$

• A third order polynomial is given by:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \beta_3 X_i^3 + u_i$$

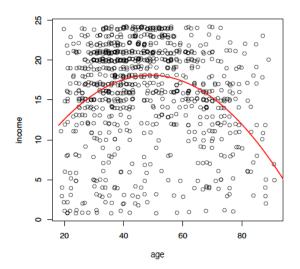
Polynomial example: Income and Age

- Classic example of a non-linear relationship between two variables: income and age
- Here we can see that simple linear regression does not seem to fit the data

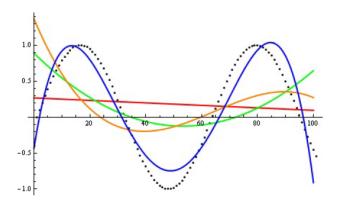


Polynomial example: Income and Age

• A second order polynomial in age fits the data a lot better:



Higher order polynomials



Covariate Adjustments in Experimental Designs

Why include PRETREATMENT covariates in experimental data analysis?

- In experiments: Regression analysis with and without PRETREATMENT covariate adjustments is unbiased!
 - Random assignment to conditions creates two groups that are, in expectation, identical prior to treatment. This implies pretreatment covariates are not correlated with the treatment
- So why adjust for covariates? Controlling for pretreatment covariates:
 - Increases precision
 - Increases statistical power
- Make sure to pre-register your model!

Including covariates in regression analysis

Let's open R studio!

Any drawbacks to adjusting for pretreatment covariates?

- In small sample (n < 20): Including pretreatment covariates introduces bias in the estimate of the ATE
- In larger samples (n > 20): Bias becomes negligible
- If you are going to include more than one pretreatment covariates, make sure sample is large enough!
 - How? Use Monte-Carlo simulation...

Including covariates in regression analysis

- Key assumption: Covariates are unaffected by treatment assignment
 - i.e., the schedule of covariates is FIXED regardless of whether an individual was assigned to treatment vs. control condition
- IMPLICATION: DO NOT adjust for posttreatment covariates!
 - It will bias ATF
 - It will turn your experimental study into a correlational study
 - It will not allow you to make any causal claims

What is a posttreatment covariate?

 Any variables that: were measured after the treatment / may be affected by the treatment

Why you SHOULD NOT adjust for posttreatment covariate?

- Adjusting for posttreatment covariates introduces bias in your estimate of the average treatment effect, even if treatment was randomly assigned to participants
 - That is because variables that are affected by treatment cannot possibly be independent of treatment assignment
- Pervasive issue in social psychology —so called mediation analysis!

Why you SHOULD NOT adjust for posttreatment covariate?

Back to R Studio!